

# Machine learning and portfolio selections. II.

László (Laci) Györfi<sup>1</sup>

<sup>1</sup>Department of Computer Science and Information Theory  
Budapest University of Technology and Economics  
Budapest, Hungary

September 22, 2007

e-mail: [gyorfi@szit.bme.hu](mailto:gyorfi@szit.bme.hu)  
[www.szit.bme.hu/~gyorfi](http://www.szit.bme.hu/~gyorfi)  
[www.szit.bme.hu/~oti/portfolio](http://www.szit.bme.hu/~oti/portfolio)

# Dynamic portfolio selection: general case

$\mathbf{x}_i = (x_i^{(1)}, \dots, x_i^{(d)})$  the return vector on day  $i$   
 $\mathbf{b} = \mathbf{b}_1$  is the portfolio vector for the first day  
initial capital  $S_0$

$$S_1 = S_0 \cdot \langle \mathbf{b}_1, \mathbf{x}_1 \rangle$$

# Dynamic portfolio selection: general case

$\mathbf{x}_i = (x_i^{(1)}, \dots, x_i^{(d)})$  the return vector on day  $i$   
 $\mathbf{b} = \mathbf{b}_1$  is the portfolio vector for the first day  
initial capital  $S_0$

$$S_1 = S_0 \cdot \langle \mathbf{b}_1, \mathbf{x}_1 \rangle$$

for the second day,  $S_1$  new initial capital, the portfolio vector  
 $\mathbf{b}_2 = \mathbf{b}(\mathbf{x}_1)$

$$S_2 = S_0 \cdot \langle \mathbf{b}_1, \mathbf{x}_1 \rangle \cdot \langle \mathbf{b}(\mathbf{x}_1), \mathbf{x}_2 \rangle .$$

# Dynamic portfolio selection: general case

$\mathbf{x}_i = (x_i^{(1)}, \dots, x_i^{(d)})$  the return vector on day  $i$   
 $\mathbf{b} = \mathbf{b}_1$  is the portfolio vector for the first day  
initial capital  $S_0$

$$S_1 = S_0 \cdot \langle \mathbf{b}_1, \mathbf{x}_1 \rangle$$

for the second day,  $S_1$  new initial capital, the portfolio vector  
 $\mathbf{b}_2 = \mathbf{b}(\mathbf{x}_1)$

$$S_2 = S_0 \cdot \langle \mathbf{b}_1, \mathbf{x}_1 \rangle \cdot \langle \mathbf{b}(\mathbf{x}_1), \mathbf{x}_2 \rangle.$$

$n$ th day a portfolio strategy  $\mathbf{b}_n = \mathbf{b}(\mathbf{x}_1, \dots, \mathbf{x}_{n-1}) = \mathbf{b}(\mathbf{x}_1^{n-1})$

$$S_n = S_0 \prod_{i=1}^n \langle \mathbf{b}(\mathbf{x}_1^{i-1}), \mathbf{x}_i \rangle = S_0 e^{nW_n(\mathbf{B})}$$

with the average growth rate

$$W_n(\mathbf{B}) = \frac{1}{n} \sum_{i=1}^n \ln \langle \mathbf{b}(\mathbf{x}_1^{i-1}), \mathbf{x}_i \rangle.$$

$\mathbf{X}_1, \mathbf{X}_2, \dots$  drawn from the vector valued stationary and ergodic process

log-optimum portfolio  $\mathbf{B}^* = \{\mathbf{b}^*(\cdot)\}$

$$\mathbf{E}\{\ln \langle \mathbf{b}^*(\mathbf{X}_1^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_1^{n-1}\} = \max_{\mathbf{b}(\cdot)} \mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{X}_1^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_1^{n-1}\}$$

$$\mathbf{X}_1^{n-1} = \mathbf{X}_1, \dots, \mathbf{X}_{n-1}$$

Algoet and Cover (1988): If  $S_n^* = S_n(\mathbf{B}^*)$  denotes the capital after day  $n$  achieved by a log-optimum portfolio strategy  $\mathbf{B}^*$ , then for any portfolio strategy  $\mathbf{B}$  with capital  $S_n = S_n(\mathbf{B})$  and for any process  $\{\mathbf{X}_n\}_{-\infty}^{\infty}$ ,

$$\limsup_{n \rightarrow \infty} \left( \frac{1}{n} \ln S_n - \frac{1}{n} \ln S_n^* \right) \leq 0 \quad \text{almost surely}$$

Algoet and Cover (1988): If  $S_n^* = S_n(\mathbf{B}^*)$  denotes the capital after day  $n$  achieved by a log-optimum portfolio strategy  $\mathbf{B}^*$ , then for any portfolio strategy  $\mathbf{B}$  with capital  $S_n = S_n(\mathbf{B})$  and for any process  $\{\mathbf{X}_n\}_{-\infty}^{\infty}$ ,

$$\limsup_{n \rightarrow \infty} \left( \frac{1}{n} \ln S_n - \frac{1}{n} \ln S_n^* \right) \leq 0 \quad \text{almost surely}$$

for stationary ergodic process  $\{\mathbf{X}_n\}_{-\infty}^{\infty}$ ,

$$\lim_{n \rightarrow \infty} \frac{1}{n} \ln S_n^* = W^* \quad \text{almost surely,}$$

where

$$W^* = \mathbf{E} \left\{ \max_{\mathbf{b}(\cdot)} \mathbf{E} \{ \ln \langle \mathbf{b}(\mathbf{X}_{-\infty}^{-1}), \mathbf{X}_0 \rangle \mid \mathbf{X}_{-\infty}^{-1} \} \right\}$$

is the maximal growth rate of any portfolio.

# Martingale difference sequences

for the proof of optimality we use the concept of martingale differences:

## Definition

there are two sequences of random variables:

$$\{Z_n\} \quad \{X_n\}$$

- $Z_n$  is a function of  $X_1, \dots, X_n$ ,
- $\mathbf{E}\{Z_n \mid X_1, \dots, X_{n-1}\} = 0$  almost surely.

Then  $\{Z_n\}$  is called martingale difference sequence with respect to  $\{X_n\}$ .



# A strong law of large numbers

**Chow Theorem:** If  $\{Z_n\}$  is a martingale difference sequence with respect to  $\{X_n\}$  and

$$\sum_{n=1}^{\infty} \frac{\mathbf{E}\{Z_n^2\}}{n^2} < \infty$$

then

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n Z_i = 0 \text{ a.s.}$$

# A weak law of large numbers

**Lemma:** If  $\{Z_n\}$  is a martingale difference sequence with respect to  $\{X_n\}$  then  $\{Z_n\}$  are uncorrelated.

# A weak law of large numbers

**Lemma:** If  $\{Z_n\}$  is a martingale difference sequence with respect to  $\{X_n\}$  then  $\{Z_n\}$  are uncorrelated.

**Proof.** Put  $i < j$ .

$$\mathbf{E}\{Z_i Z_j\}$$

# A weak law of large numbers

**Lemma:** If  $\{Z_n\}$  is a martingale difference sequence with respect to  $\{X_n\}$  then  $\{Z_n\}$  are uncorrelated.

**Proof.** Put  $i < j$ .

$$\mathbf{E}\{Z_i Z_j\} = \mathbf{E}\{\mathbf{E}\{Z_i Z_j \mid X_1, \dots, X_{j-1}\}\}$$

# A weak law of large numbers

**Lemma:** If  $\{Z_n\}$  is a martingale difference sequence with respect to  $\{X_n\}$  then  $\{Z_n\}$  are uncorrelated.

**Proof.** Put  $i < j$ .

$$\begin{aligned}\mathbf{E}\{Z_i Z_j\} &= \mathbf{E}\{\mathbf{E}\{Z_i Z_j \mid X_1, \dots, X_{j-1}\}\} \\ &= \mathbf{E}\{Z_i \mathbf{E}\{Z_j \mid X_1, \dots, X_{j-1}\}\}\end{aligned}$$

# A weak law of large numbers

**Lemma:** If  $\{Z_n\}$  is a martingale difference sequence with respect to  $\{X_n\}$  then  $\{Z_n\}$  are uncorrelated.

**Proof.** Put  $i < j$ .

$$\begin{aligned}\mathbf{E}\{Z_i Z_j\} &= \mathbf{E}\{\mathbf{E}\{Z_i Z_j \mid X_1, \dots, X_{j-1}\}\} \\ &= \mathbf{E}\{Z_i \mathbf{E}\{Z_j \mid X_1, \dots, X_{j-1}\}\} \\ &= \mathbf{E}\{Z_i \cdot 0\}\end{aligned}$$

# A weak law of large numbers

**Lemma:** If  $\{Z_n\}$  is a martingale difference sequence with respect to  $\{X_n\}$  then  $\{Z_n\}$  are uncorrelated.

**Proof.** Put  $i < j$ .

$$\begin{aligned}\mathbf{E}\{Z_i Z_j\} &= \mathbf{E}\{\mathbf{E}\{Z_i Z_j \mid X_1, \dots, X_{j-1}\}\} \\ &= \mathbf{E}\{Z_i \mathbf{E}\{Z_j \mid X_1, \dots, X_{j-1}\}\} \\ &= \mathbf{E}\{Z_i \cdot 0\} = 0\end{aligned}$$

# A weak law of large numbers

**Lemma:** If  $\{Z_n\}$  is a martingale difference sequence with respect to  $\{X_n\}$  then  $\{Z_n\}$  are uncorrelated.

**Proof.** Put  $i < j$ .

$$\begin{aligned}\mathbf{E}\{Z_i Z_j\} &= \mathbf{E}\{\mathbf{E}\{Z_i Z_j \mid X_1, \dots, X_{j-1}\}\} \\ &= \mathbf{E}\{Z_i \mathbf{E}\{Z_j \mid X_1, \dots, X_{j-1}\}\} \\ &= \mathbf{E}\{Z_i \cdot 0\} = 0\end{aligned}$$

**Corollary**

$$\mathbf{E} \left\{ \left( \frac{1}{n} \sum_{i=1}^n Z_i \right)^2 \right\}$$



# A weak law of large numbers

**Lemma:** If  $\{Z_n\}$  is a martingale difference sequence with respect to  $\{X_n\}$  then  $\{Z_n\}$  are uncorrelated.

**Proof.** Put  $i < j$ .

$$\begin{aligned}\mathbf{E}\{Z_i Z_j\} &= \mathbf{E}\{\mathbf{E}\{Z_i Z_j \mid X_1, \dots, X_{j-1}\}\} \\ &= \mathbf{E}\{Z_i \mathbf{E}\{Z_j \mid X_1, \dots, X_{j-1}\}\} \\ &= \mathbf{E}\{Z_i \cdot 0\} = 0\end{aligned}$$

**Corollary**

$$\mathbf{E}\left\{\left(\frac{1}{n} \sum_{i=1}^n Z_i\right)^2\right\} = \frac{1}{n^2} \sum_{i=1}^n \sum_{j=1}^n \mathbf{E}\{Z_i Z_j\}$$

# A weak law of large numbers

**Lemma:** If  $\{Z_n\}$  is a martingale difference sequence with respect to  $\{X_n\}$  then  $\{Z_n\}$  are uncorrelated.

**Proof.** Put  $i < j$ .

$$\begin{aligned}\mathbf{E}\{Z_i Z_j\} &= \mathbf{E}\{\mathbf{E}\{Z_i Z_j \mid X_1, \dots, X_{j-1}\}\} \\ &= \mathbf{E}\{Z_i \mathbf{E}\{Z_j \mid X_1, \dots, X_{j-1}\}\} \\ &= \mathbf{E}\{Z_i \cdot 0\} = 0\end{aligned}$$

**Corollary**

$$\begin{aligned}\mathbf{E}\left\{\left(\frac{1}{n} \sum_{i=1}^n Z_i\right)^2\right\} &= \frac{1}{n^2} \sum_{i=1}^n \sum_{j=1}^n \mathbf{E}\{Z_i Z_j\} \\ &= \frac{1}{n^2} \sum_{i=1}^n \mathbf{E}\{Z_i^2\}\end{aligned}$$

# A weak law of large numbers

**Lemma:** If  $\{Z_n\}$  is a martingale difference sequence with respect to  $\{X_n\}$  then  $\{Z_n\}$  are uncorrelated.

**Proof.** Put  $i < j$ .

$$\begin{aligned}\mathbf{E}\{Z_i Z_j\} &= \mathbf{E}\{\mathbf{E}\{Z_i Z_j \mid X_1, \dots, X_{j-1}\}\} \\ &= \mathbf{E}\{Z_i \mathbf{E}\{Z_j \mid X_1, \dots, X_{j-1}\}\} \\ &= \mathbf{E}\{Z_i \cdot 0\} = 0\end{aligned}$$

**Corollary**

$$\begin{aligned}\mathbf{E}\left\{\left(\frac{1}{n} \sum_{i=1}^n Z_i\right)^2\right\} &= \frac{1}{n^2} \sum_{i=1}^n \sum_{j=1}^n \mathbf{E}\{Z_i Z_j\} \\ &= \frac{1}{n^2} \sum_{i=1}^n \mathbf{E}\{Z_i^2\} \\ &\rightarrow 0\end{aligned}$$

if, for example,  $\mathbf{E}\{Z_i^2\}$  is a bounded sequence.

# Constructing martingale difference sequence

$\{Y_n\}$  is an arbitrary sequence such that  $Y_n$  is a function of  $X_1, \dots, X_n$

# Constructing martingale difference sequence

$\{Y_n\}$  is an arbitrary sequence such that  $Y_n$  is a function of  $X_1, \dots, X_n$

Put

$$Z_n = Y_n - \mathbf{E}\{Y_n \mid X_1, \dots, X_{n-1}\}$$

Then  $\{Z_n\}$  is a martingale difference sequence:

# Constructing martingale difference sequence

$\{Y_n\}$  is an arbitrary sequence such that  $Y_n$  is a function of  $X_1, \dots, X_n$

Put

$$Z_n = Y_n - \mathbf{E}\{Y_n \mid X_1, \dots, X_{n-1}\}$$

Then  $\{Z_n\}$  is a martingale difference sequence:

- $Z_n$  is a function of  $X_1, \dots, X_n$ ,

# Constructing martingale difference sequence

$\{Y_n\}$  is an arbitrary sequence such that  $Y_n$  is a function of  $X_1, \dots, X_n$

Put

$$Z_n = Y_n - \mathbf{E}\{Y_n \mid X_1, \dots, X_{n-1}\}$$

Then  $\{Z_n\}$  is a martingale difference sequence:

- $Z_n$  is a function of  $X_1, \dots, X_n$ ,
- 

$$\mathbf{E}\{Z_n \mid X_1, \dots, X_{n-1}\}$$

# Constructing martingale difference sequence

$\{Y_n\}$  is an arbitrary sequence such that  $Y_n$  is a function of  $X_1, \dots, X_n$

Put

$$Z_n = Y_n - \mathbf{E}\{Y_n \mid X_1, \dots, X_{n-1}\}$$

Then  $\{Z_n\}$  is a martingale difference sequence:

- $Z_n$  is a function of  $X_1, \dots, X_n$ ,
- 

$$\begin{aligned} & \mathbf{E}\{Z_n \mid X_1, \dots, X_{n-1}\} \\ &= \mathbf{E}\{Y_n - \mathbf{E}\{Y_n \mid X_1, \dots, X_{n-1}\} \mid X_1, \dots, X_{n-1}\} \\ &= 0 \end{aligned}$$

almost surely.



log-optimum portfolio  $\mathbf{B}^* = \{\mathbf{b}^*(\cdot)\}$

$$\mathbf{E}\{\ln \langle \mathbf{b}^*(\mathbf{X}_1^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_1^{n-1}\} = \max_{\mathbf{b}(\cdot)} \mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{X}_1^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_1^{n-1}\}$$

log-optimum portfolio  $\mathbf{B}^* = \{\mathbf{b}^*(\cdot)\}$

$$\mathbf{E}\{\ln \langle \mathbf{b}^*(\mathbf{X}_1^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_1^{n-1}\} = \max_{\mathbf{b}(\cdot)} \mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{X}_1^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_1^{n-1}\}$$

If  $S_n^* = S_n(\mathbf{B}^*)$  denotes the capital after day  $n$  achieved by a log-optimum portfolio strategy  $\mathbf{B}^*$ , then for any portfolio strategy  $\mathbf{B}$  with capital  $S_n = S_n(\mathbf{B})$  and for any process  $\{\mathbf{X}_n\}_{-\infty}^{\infty}$ ,

$$\limsup_{n \rightarrow \infty} \left( \frac{1}{n} \ln S_n - \frac{1}{n} \ln S_n^* \right) \leq 0 \quad \text{almost surely}$$

# Proof of optimality

$$\frac{1}{n} \ln S_n = \frac{1}{n} \sum_{i=1}^n \ln \langle \mathbf{b}(\mathbf{X}_1^{i-1}), \mathbf{X}_i \rangle$$

# Proof of optimality

$$\begin{aligned}\frac{1}{n} \ln S_n &= \frac{1}{n} \sum_{i=1}^n \ln \langle \mathbf{b}(\mathbf{X}_1^{i-1}), \mathbf{X}_i \rangle \\ &= \frac{1}{n} \sum_{i=1}^n \mathbf{E} \{ \ln \langle \mathbf{b}(\mathbf{X}_1^{i-1}), \mathbf{X}_i \rangle \mid \mathbf{X}_1^{i-1} \} \\ &+ \frac{1}{n} \sum_{i=1}^n \left( \ln \langle \mathbf{b}(\mathbf{X}_1^{i-1}), \mathbf{X}_i \rangle - \mathbf{E} \{ \ln \langle \mathbf{b}(\mathbf{X}_1^{i-1}), \mathbf{X}_i \rangle \mid \mathbf{X}_1^{i-1} \} \right)\end{aligned}$$

# Proof of optimality

$$\begin{aligned}\frac{1}{n} \ln S_n &= \frac{1}{n} \sum_{i=1}^n \ln \langle \mathbf{b}(\mathbf{X}_1^{i-1}), \mathbf{X}_i \rangle \\ &= \frac{1}{n} \sum_{i=1}^n \mathbf{E} \{ \ln \langle \mathbf{b}(\mathbf{X}_1^{i-1}), \mathbf{X}_i \rangle \mid \mathbf{X}_1^{i-1} \} \\ &+ \frac{1}{n} \sum_{i=1}^n \left( \ln \langle \mathbf{b}(\mathbf{X}_1^{i-1}), \mathbf{X}_i \rangle - \mathbf{E} \{ \ln \langle \mathbf{b}(\mathbf{X}_1^{i-1}), \mathbf{X}_i \rangle \mid \mathbf{X}_1^{i-1} \} \right)\end{aligned}$$

and

$$\begin{aligned}\frac{1}{n} \ln S_n^* &= \frac{1}{n} \sum_{i=1}^n \mathbf{E} \{ \ln \langle \mathbf{b}^*(\mathbf{X}_1^{i-1}), \mathbf{X}_i \rangle \mid \mathbf{X}_1^{i-1} \} \\ &+ \frac{1}{n} \sum_{i=1}^n \left( \ln \langle \mathbf{b}^*(\mathbf{X}_1^{i-1}), \mathbf{X}_i \rangle - \mathbf{E} \{ \ln \langle \mathbf{b}^*(\mathbf{X}_1^{i-1}), \mathbf{X}_i \rangle \mid \mathbf{X}_1^{i-1} \} \right)\end{aligned}$$

These limit relations give rise to the following definition:

## Definition

An empirical (data driven) portfolio strategy  $\mathbf{B}$  is called **universally consistent with respect to a class  $\mathcal{C}$  of stationary and ergodic processes  $\{\mathbf{X}_n\}_{-\infty}^{\infty}$** , if for each process in the class,

$$\lim_{n \rightarrow \infty} \frac{1}{n} \ln S_n(\mathbf{B}) = W^* \quad \text{almost surely.}$$

$$\mathbf{E}\{\ln \langle \mathbf{b}^*(\mathbf{X}_1^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_1^{n-1}\} = \max_{\mathbf{b}(\cdot)} \mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{X}_1^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_1^{n-1}\}$$

$$\mathbf{E}\{\ln \langle \mathbf{b}^*(\mathbf{X}_1^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_1^{n-1}\} = \max_{\mathbf{b}(\cdot)} \mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{X}_1^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_1^{n-1}\}$$

fixed integer  $k > 0$

$$\mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{X}_1^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_1^{n-1}\} \approx \mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{X}_{n-k}^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_{n-k}^{n-1}\}$$



$$\mathbf{E}\{\ln \langle \mathbf{b}^*(\mathbf{X}_1^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_1^{n-1}\} = \max_{\mathbf{b}(\cdot)} \mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{X}_1^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_1^{n-1}\}$$

fixed integer  $k > 0$

$$\mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{X}_1^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_1^{n-1}\} \approx \mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{X}_{n-k}^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_{n-k}^{n-1}\}$$

and

$$\mathbf{b}^*(\mathbf{X}_1^{n-1}) \approx \mathbf{b}_k(\mathbf{X}_{n-k}^{n-1}) = \arg \max_{\mathbf{b}(\cdot)} \mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{X}_{n-k}^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_{n-k}^{n-1}\}$$

$$\mathbf{E}\{\ln \langle \mathbf{b}^*(\mathbf{X}_1^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_1^{n-1}\} = \max_{\mathbf{b}(\cdot)} \mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{X}_1^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_1^{n-1}\}$$

fixed integer  $k > 0$

$$\mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{X}_1^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_1^{n-1}\} \approx \mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{X}_{n-k}^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_{n-k}^{n-1}\}$$

and

$$\mathbf{b}^*(\mathbf{X}_1^{n-1}) \approx \mathbf{b}_k(\mathbf{X}_{n-k}^{n-1}) = \arg \max_{\mathbf{b}(\cdot)} \mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{X}_{n-k}^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_{n-k}^{n-1}\}$$

because of stationarity

$$\begin{aligned} \mathbf{b}_k(\mathbf{x}_1^k) &= \arg \max_{\mathbf{b}(\cdot)} \mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{x}_1^k), \mathbf{X}_{k+1} \rangle \mid \mathbf{X}_1^k = \mathbf{x}_1^k\} \\ &= \arg \max_{\mathbf{b}} \mathbf{E}\{\ln \langle \mathbf{b}, \mathbf{X}_{k+1} \rangle \mid \mathbf{X}_1^k = \mathbf{x}_1^k\}, \end{aligned}$$

$$\mathbf{E}\{\ln \langle \mathbf{b}^*(\mathbf{X}_1^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_1^{n-1}\} = \max_{\mathbf{b}(\cdot)} \mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{X}_1^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_1^{n-1}\}$$

fixed integer  $k > 0$

$$\mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{X}_1^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_1^{n-1}\} \approx \mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{X}_{n-k}^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_{n-k}^{n-1}\}$$

and

$$\mathbf{b}^*(\mathbf{X}_1^{n-1}) \approx \mathbf{b}_k(\mathbf{X}_{n-k}^{n-1}) = \arg \max_{\mathbf{b}(\cdot)} \mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{X}_{n-k}^{n-1}), \mathbf{X}_n \rangle \mid \mathbf{X}_{n-k}^{n-1}\}$$

because of stationarity

$$\begin{aligned} \mathbf{b}_k(\mathbf{x}_1^k) &= \arg \max_{\mathbf{b}(\cdot)} \mathbf{E}\{\ln \langle \mathbf{b}(\mathbf{x}_1^k), \mathbf{X}_{k+1} \rangle \mid \mathbf{X}_1^k = \mathbf{x}_1^k\} \\ &= \arg \max_{\mathbf{b}} \mathbf{E}\{\ln \langle \mathbf{b}, \mathbf{X}_{k+1} \rangle \mid \mathbf{X}_1^k = \mathbf{x}_1^k\}, \end{aligned}$$

which is the maximization of the regression function

$$m_{\mathbf{b}}(\mathbf{x}_1^k) = \mathbf{E}\{\ln \langle \mathbf{b}, \mathbf{X}_{k+1} \rangle \mid \mathbf{X}_1^k = \mathbf{x}_1^k\}$$

# Regression function

$Y$  real valued

$X$  observation vector

# Regression function

$Y$  real valued

$X$  observation vector

Regression function

$$m(x) = \mathbf{E}\{Y \mid X = x\}$$

i.i.d. data:  $D_n = \{(X_1, Y_1), \dots, (X_n, Y_n)\}$

# Regression function

$Y$  real valued

$X$  observation vector

Regression function

$$m(x) = \mathbf{E}\{Y \mid X = x\}$$

i.i.d. data:  $D_n = \{(X_1, Y_1), \dots, (X_n, Y_n)\}$

Regression function estimate

$$m_n(x) = m_n(x, D_n)$$

# Regression function

$Y$  real valued

$X$  observation vector

Regression function

$$m(x) = \mathbf{E}\{Y \mid X = x\}$$

i.i.d. data:  $D_n = \{(X_1, Y_1), \dots, (X_n, Y_n)\}$

Regression function estimate

$$m_n(x) = m_n(x, D_n)$$

local averaging estimates

$$m_n(x) = \sum_{i=1}^n W_{ni}(x; X_1, \dots, X_n) Y_i$$

# Regression function

$Y$  real valued

$X$  observation vector

Regression function

$$m(x) = \mathbf{E}\{Y \mid X = x\}$$

i.i.d. data:  $D_n = \{(X_1, Y_1), \dots, (X_n, Y_n)\}$

Regression function estimate

$$m_n(x) = m_n(x, D_n)$$

local averaging estimates

$$m_n(x) = \sum_{i=1}^n W_{ni}(x; X_1, \dots, X_n) Y_i$$

L. Györfi, M. Kohler, A. Krzyzak, H. Walk (2002) *A Distribution-Free Theory of Nonparametric Regression*, Springer-Verlag, New York.



$$X \sim \mathbf{x}_1^k$$

$$\begin{aligned} X &\sim \mathbf{X}_1^k \\ Y &\sim \ln \langle \mathbf{b}, \mathbf{X}_{k+1} \rangle \end{aligned}$$

$$X \sim \mathbf{X}_1^k$$

$$Y \sim \ln \langle \mathbf{b}, \mathbf{X}_{k+1} \rangle$$

$$m(x) = \mathbf{E}\{Y \mid X = x\} \sim m_{\mathbf{b}}(\mathbf{x}_1^k) = \mathbf{E}\{\ln \langle \mathbf{b}, \mathbf{X}_{k+1} \rangle \mid \mathbf{X}_1^k = \mathbf{x}_1^k\}$$

# Partitioning regression estimate

Partition  $\mathcal{P}_n = \{A_{n,1}, A_{n,2} \dots\}$

# Partitioning regression estimate

Partition  $\mathcal{P}_n = \{A_{n,1}, A_{n,2}, \dots\}$

$A_n(x)$  is the cell of the partition  $\mathcal{P}_n$  into which  $x$  falls

# Partitioning regression estimate

Partition  $\mathcal{P}_n = \{A_{n,1}, A_{n,2}, \dots\}$

$A_n(x)$  is the cell of the partition  $\mathcal{P}_n$  into which  $x$  falls

$$m_n(x) = \frac{\sum_{i=1}^n Y_i I_{[X_i \in A_n(x)]}}{\sum_{i=1}^n I_{[X_i \in A_n(x)]}}$$

# Partitioning regression estimate

Partition  $\mathcal{P}_n = \{A_{n,1}, A_{n,2}, \dots\}$

$A_n(x)$  is the cell of the partition  $\mathcal{P}_n$  into which  $x$  falls

$$m_n(x) = \frac{\sum_{i=1}^n Y_i I_{[X_i \in A_n(x)]}}{\sum_{i=1}^n I_{[X_i \in A_n(x)]}}$$

Let  $G_n$  be the quantizer corresponding to the partition  $\mathcal{P}_n$ :

$G_n(x) = j$  if  $x \in A_{n,j}$ .

# Partitioning regression estimate

Partition  $\mathcal{P}_n = \{A_{n,1}, A_{n,2}, \dots\}$

$A_n(x)$  is the cell of the partition  $\mathcal{P}_n$  into which  $x$  falls

$$m_n(x) = \frac{\sum_{i=1}^n Y_i I_{[X_i \in A_n(x)]}}{\sum_{i=1}^n I_{[X_i \in A_n(x)]}}$$

Let  $G_n$  be the quantizer corresponding to the partition  $\mathcal{P}_n$ :

$G_n(x) = j$  if  $x \in A_{n,j}$ .

the set of matches

$$I_n(x) = \{i \leq n : G_n(x) = G_n(X_i)\}$$

Then

$$m_n(x) = \frac{\sum_{i \in I_n(x)} Y_i}{|I_n(x)|}.$$



# Partitioning-based portfolio selection

fix  $k, \ell = 1, 2, \dots$

$\mathcal{P}_\ell = \{A_{\ell,j}, j = 1, 2, \dots, m_\ell\}$  finite partitions of  $\mathbb{R}^d$ ,

# Partitioning-based portfolio selection

fix  $k, \ell = 1, 2, \dots$

$\mathcal{P}_\ell = \{A_{\ell,j}, j = 1, 2, \dots, m_\ell\}$  finite partitions of  $\mathbb{R}^d$ ,

$G_\ell$  be the corresponding quantizer:  $G_\ell(\mathbf{x}) = j$ , if  $\mathbf{x} \in A_{\ell,j}$ .

# Partitioning-based portfolio selection

fix  $k, \ell = 1, 2, \dots$

$\mathcal{P}_\ell = \{A_{\ell,j}, j = 1, 2, \dots, m_\ell\}$  finite partitions of  $\mathbb{R}^d$ ,

$G_\ell$  be the corresponding quantizer:  $G_\ell(\mathbf{x}) = j$ , if  $\mathbf{x} \in A_{\ell,j}$ .

$G_\ell(\mathbf{x}_1^n) = G_\ell(\mathbf{x}_1), \dots, G_\ell(\mathbf{x}_n)$ ,

# Partitioning-based portfolio selection

fix  $k, \ell = 1, 2, \dots$

$\mathcal{P}_\ell = \{A_{\ell,j}, j = 1, 2, \dots, m_\ell\}$  finite partitions of  $\mathbb{R}^d$ ,

$G_\ell$  be the corresponding quantizer:  $G_\ell(\mathbf{x}) = j$ , if  $\mathbf{x} \in A_{\ell,j}$ .

$G_\ell(\mathbf{x}_1^n) = G_\ell(\mathbf{x}_1), \dots, G_\ell(\mathbf{x}_n)$ ,

the set of matches:

$$J_n = \{k < i < n : G_\ell(\mathbf{x}_{i-k}^{i-1}) = G_\ell(\mathbf{x}_{n-k}^{n-1})\}$$

# Partitioning-based portfolio selection

fix  $k, \ell = 1, 2, \dots$

$\mathcal{P}_\ell = \{A_{\ell,j}, j = 1, 2, \dots, m_\ell\}$  finite partitions of  $\mathbb{R}^d$ ,

$G_\ell$  be the corresponding quantizer:  $G_\ell(\mathbf{x}) = j$ , if  $\mathbf{x} \in A_{\ell,j}$ .

$G_\ell(\mathbf{x}_1^n) = G_\ell(\mathbf{x}_1), \dots, G_\ell(\mathbf{x}_n)$ ,

the set of matches:

$$J_n = \{k < i < n : G_\ell(\mathbf{x}_{i-k}^{i-1}) = G_\ell(\mathbf{x}_{n-k}^{n-1})\}$$

$$\mathbf{b}^{(k,\ell)}(\mathbf{x}_1^{n-1}) = \arg \max_{\mathbf{b}} \sum_{i \in J_n} \ln \langle \mathbf{b}, \mathbf{x}_i \rangle$$

if the set  $J_n$  is non-void, and  $\mathbf{b}_0 = (1/d, \dots, 1/d)$  otherwise.

for fixed  $k, \ell = 1, 2, \dots,$

$\mathbf{B}^{(k,\ell)} = \{\mathbf{b}^{(k,\ell)}(\cdot)\}$ , are called elementary portfolios

for fixed  $k, \ell = 1, 2, \dots,$

$\mathbf{B}^{(k,\ell)} = \{\mathbf{b}^{(k,\ell)}(\cdot)\}$ , are called elementary portfolios

That is,  $\mathbf{b}_n^{(k,\ell)}$  quantizes the sequence  $\mathbf{x}_1^{n-1}$  according to the partition  $\mathcal{P}_\ell$ , and browses through all past appearances of the last seen quantized string  $G_\ell(\mathbf{x}_{n-k}^{n-1})$  of length  $k$ .

Then it designs a fixed portfolio vector according to the returns on the days following the occurrence of the string.

# Combining elementary portfolios

How to choose  $k, \ell$

- small  $k$  or small  $\ell$ : large bias
- large  $k$  and large  $\ell$ : few matching, large variance



How to choose  $k, \ell$

- small  $k$  or small  $\ell$ : large bias
- large  $k$  and large  $\ell$ : few matching, large variance

Machine learning: combination of experts

N. Cesa-Bianchi and G. Lugosi, *Prediction, Learning, and Games*.  
Cambridge University Press, 2006.

# Exponential weighing

combine the elementary portfolio strategies  $\mathbf{B}^{(k,\ell)} = \{\mathbf{b}_n^{(k,\ell)}\}$

# Exponential weighing

combine the elementary portfolio strategies  $\mathbf{B}^{(k,\ell)} = \{\mathbf{b}_n^{(k,\ell)}\}$   
let  $\{q_{k,\ell}\}$  be a probability distribution on the set of all pairs  $(k, \ell)$   
such that for all  $k, \ell$ ,  $q_{k,\ell} > 0$ .

# Exponential weighing

combine the elementary portfolio strategies  $\mathbf{B}^{(k,\ell)} = \{\mathbf{b}_n^{(k,\ell)}\}$

let  $\{q_{k,\ell}\}$  be a probability distribution on the set of all pairs  $(k, \ell)$

such that for all  $k, \ell$ ,  $q_{k,\ell} > 0$ .

for  $\eta > 0$  put

$$w_{n,k,\ell} = q_{k,\ell} e^{\eta \ln S_{n-1}(\mathbf{B}^{(k,\ell)})}$$

# Exponential weighing

combine the elementary portfolio strategies  $\mathbf{B}^{(k,\ell)} = \{\mathbf{b}_n^{(k,\ell)}\}$

let  $\{q_{k,\ell}\}$  be a probability distribution on the set of all pairs  $(k, \ell)$  such that for all  $k, \ell$ ,  $q_{k,\ell} > 0$ .

for  $\eta > 0$  put

$$w_{n,k,\ell} = q_{k,\ell} e^{\eta \ln S_{n-1}(\mathbf{B}^{(k,\ell)})}$$

for  $\eta = 1$ ,

$$w_{n,k,\ell} = q_{k,\ell} e^{\ln S_{n-1}(\mathbf{B}^{(k,\ell)})} = q_{k,\ell} S_{n-1}(\mathbf{B}^{(k,\ell)})$$

# Exponential weighing

combine the elementary portfolio strategies  $\mathbf{B}^{(k,\ell)} = \{\mathbf{b}_n^{(k,\ell)}\}$   
let  $\{q_{k,\ell}\}$  be a probability distribution on the set of all pairs  $(k, \ell)$   
such that for all  $k, \ell$ ,  $q_{k,\ell} > 0$ .

for  $\eta > 0$  put

$$w_{n,k,\ell} = q_{k,\ell} e^{\eta \ln S_{n-1}(\mathbf{B}^{(k,\ell)})}$$

for  $\eta = 1$ ,

$$w_{n,k,\ell} = q_{k,\ell} e^{\ln S_{n-1}(\mathbf{B}^{(k,\ell)})} = q_{k,\ell} S_{n-1}(\mathbf{B}^{(k,\ell)})$$

and

$$v_{n,k,\ell} = \frac{w_{n,k,\ell}}{\sum_{i,j} w_{n,i,j}}.$$

# Exponential weighing

combine the elementary portfolio strategies  $\mathbf{B}^{(k,\ell)} = \{\mathbf{b}_n^{(k,\ell)}\}$   
let  $\{q_{k,\ell}\}$  be a probability distribution on the set of all pairs  $(k, \ell)$   
such that for all  $k, \ell$ ,  $q_{k,\ell} > 0$ .

for  $\eta > 0$  put

$$w_{n,k,\ell} = q_{k,\ell} e^{\eta \ln S_{n-1}(\mathbf{B}^{(k,\ell)})}$$

for  $\eta = 1$ ,

$$w_{n,k,\ell} = q_{k,\ell} e^{\ln S_{n-1}(\mathbf{B}^{(k,\ell)})} = q_{k,\ell} S_{n-1}(\mathbf{B}^{(k,\ell)})$$

and

$$v_{n,k,\ell} = \frac{w_{n,k,\ell}}{\sum_{i,j} w_{n,i,j}}.$$

the combined portfolio  $\mathbf{b}$ :

$$\mathbf{b}_n(\mathbf{x}_1^{n-1}) = \sum_{k,\ell} v_{n,k,\ell} \mathbf{b}_n^{(k,\ell)}(\mathbf{x}_1^{n-1}).$$

$$S_n(\mathbf{B}) = \prod_{i=1}^n \langle \mathbf{b}_i(\mathbf{x}_1^{i-1}), \mathbf{x}_i \rangle$$



$$\begin{aligned} S_n(\mathbf{B}) &= \prod_{i=1}^n \langle \mathbf{b}_i(\mathbf{x}_1^{i-1}), \mathbf{x}_i \rangle \\ &= \prod_{i=1}^n \frac{\sum_{k,\ell} w_{i,k,\ell} \langle \mathbf{b}_i^{(k,\ell)}(\mathbf{x}_1^{i-1}), \mathbf{x}_i \rangle}{\sum_{k,\ell} w_{i,k,\ell}} \end{aligned}$$

$$\begin{aligned}
S_n(\mathbf{B}) &= \prod_{i=1}^n \langle \mathbf{b}_i(\mathbf{x}_1^{i-1}), \mathbf{x}_i \rangle \\
&= \prod_{i=1}^n \frac{\sum_{k,\ell} w_{i,k,\ell} \langle \mathbf{b}_i^{(k,\ell)}(\mathbf{x}_1^{i-1}), \mathbf{x}_i \rangle}{\sum_{k,\ell} w_{i,k,\ell}} \\
&= \prod_{i=1}^n \frac{\sum_{k,\ell} q_{k,\ell} S_{i-1}(\mathbf{B}^{(k,\ell)}) \langle \mathbf{b}_i^{(k,\ell)}(\mathbf{x}_1^{i-1}), \mathbf{x}_i \rangle}{\sum_{k,\ell} q_{k,\ell} S_{i-1}(\mathbf{B}^{(k,\ell)})}
\end{aligned}$$

$$\begin{aligned}
S_n(\mathbf{B}) &= \prod_{i=1}^n \langle \mathbf{b}_i(\mathbf{x}_1^{i-1}), \mathbf{x}_i \rangle \\
&= \prod_{i=1}^n \frac{\sum_{k,\ell} w_{i,k,\ell} \langle \mathbf{b}_i^{(k,\ell)}(\mathbf{x}_1^{i-1}), \mathbf{x}_i \rangle}{\sum_{k,\ell} w_{i,k,\ell}} \\
&= \prod_{i=1}^n \frac{\sum_{k,\ell} q_{k,\ell} S_{i-1}(\mathbf{B}^{(k,\ell)}) \langle \mathbf{b}_i^{(k,\ell)}(\mathbf{x}_1^{i-1}), \mathbf{x}_i \rangle}{\sum_{k,\ell} q_{k,\ell} S_{i-1}(\mathbf{B}^{(k,\ell)})} \\
&= \prod_{i=1}^n \frac{\sum_{k,\ell} q_{k,\ell} S_i(\mathbf{B}^{(k,\ell)})}{\sum_{k,\ell} q_{k,\ell} S_{i-1}(\mathbf{B}^{(k,\ell)})}
\end{aligned}$$

$$\begin{aligned}
S_n(\mathbf{B}) &= \prod_{i=1}^n \langle \mathbf{b}_i(\mathbf{x}_1^{i-1}), \mathbf{x}_i \rangle \\
&= \prod_{i=1}^n \frac{\sum_{k,\ell} w_{i,k,\ell} \langle \mathbf{b}_i^{(k,\ell)}(\mathbf{x}_1^{i-1}), \mathbf{x}_i \rangle}{\sum_{k,\ell} w_{i,k,\ell}} \\
&= \prod_{i=1}^n \frac{\sum_{k,\ell} q_{k,\ell} S_{i-1}(\mathbf{B}^{(k,\ell)}) \langle \mathbf{b}_i^{(k,\ell)}(\mathbf{x}_1^{i-1}), \mathbf{x}_i \rangle}{\sum_{k,\ell} q_{k,\ell} S_{i-1}(\mathbf{B}^{(k,\ell)})} \\
&= \prod_{i=1}^n \frac{\sum_{k,\ell} q_{k,\ell} S_i(\mathbf{B}^{(k,\ell)})}{\sum_{k,\ell} q_{k,\ell} S_{i-1}(\mathbf{B}^{(k,\ell)})} \\
&= \sum_{k,\ell} q_{k,\ell} S_n(\mathbf{B}^{(k,\ell)}),
\end{aligned}$$

The strategy  $\mathbf{B} = \mathbf{B}^H$  then arises from weighing the elementary portfolio strategies  $\mathbf{B}^{(k,\ell)} = \{\mathbf{b}_n^{(k,\ell)}\}$  such that the investor's capital becomes

$$S_n(\mathbf{B}) = \sum_{k,\ell} q_{k,\ell} S_n(\mathbf{B}^{(k,\ell)}).$$

Assume that

- (a) the sequence of partitions is nested, that is, any cell of  $\mathcal{P}_{\ell+1}$  is a subset of a cell of  $\mathcal{P}_{\ell}$ ,  $\ell = 1, 2, \dots$ ;
- (b) if  $\text{diam}(A) = \sup_{\mathbf{x}, \mathbf{y} \in A} \|\mathbf{x} - \mathbf{y}\|$  denotes the diameter of a set, then for any sphere  $S$  centered at the origin

$$\lim_{\ell \rightarrow \infty} \max_{j: A_{\ell, j} \cap S \neq \emptyset} \text{diam}(A_{\ell, j}) = 0 .$$

Then the portfolio scheme  $\mathbf{B}^H$  defined above is universally consistent with respect to the class of all ergodic processes such that  $\mathbf{E}\{|\ln X^{(j)}| < \infty, \text{ for } j = 1, 2, \dots, d.$

L. Györfi, D. Schäfer (2003) "Nonparametric prediction", in  
*Advances in Learning Theory: Methods, Models and Applications*,  
J. A. K. Suykens, G. Horváth, S. Basu, C. Micchelli, J. Vandevallé  
(Eds.), IOS Press, NATO Science Series, pp. 341-356.  
[www.szit.bme.hu/~gyorfi/histog.ps](http://www.szit.bme.hu/~gyorfi/histog.ps)

We have to prove that

$$\liminf_{n \rightarrow \infty} W_n(\mathbf{B}) = \liminf_{n \rightarrow \infty} \frac{1}{n} \ln S_n(\mathbf{B}) \geq W^* \quad \text{a.s.}$$

W.l.o.g. we may assume  $S_0 = 1$ , so that

$$W_n(\mathbf{B}) = \frac{1}{n} \ln S_n(\mathbf{B})$$



We have to prove that

$$\liminf_{n \rightarrow \infty} W_n(\mathbf{B}) = \liminf_{n \rightarrow \infty} \frac{1}{n} \ln S_n(\mathbf{B}) \geq W^* \quad \text{a.s.}$$

W.l.o.g. we may assume  $S_0 = 1$ , so that

$$\begin{aligned} W_n(\mathbf{B}) &= \frac{1}{n} \ln S_n(\mathbf{B}) \\ &= \frac{1}{n} \ln \left( \sum_{k,\ell} q_{k,\ell} S_n(\mathbf{B}^{(k,\ell)}) \right) \end{aligned}$$

We have to prove that

$$\liminf_{n \rightarrow \infty} W_n(\mathbf{B}) = \liminf_{n \rightarrow \infty} \frac{1}{n} \ln S_n(\mathbf{B}) \geq W^* \quad \text{a.s.}$$

W.l.o.g. we may assume  $S_0 = 1$ , so that

$$\begin{aligned} W_n(\mathbf{B}) &= \frac{1}{n} \ln S_n(\mathbf{B}) \\ &= \frac{1}{n} \ln \left( \sum_{k,\ell} q_{k,\ell} S_n(\mathbf{B}^{(k,\ell)}) \right) \\ &\geq \frac{1}{n} \ln \left( \sup_{k,\ell} q_{k,\ell} S_n(\mathbf{B}^{(k,\ell)}) \right) \end{aligned}$$

We have to prove that

$$\liminf_{n \rightarrow \infty} W_n(\mathbf{B}) = \liminf_{n \rightarrow \infty} \frac{1}{n} \ln S_n(\mathbf{B}) \geq W^* \quad \text{a.s.}$$

W.l.o.g. we may assume  $S_0 = 1$ , so that

$$\begin{aligned} W_n(\mathbf{B}) &= \frac{1}{n} \ln S_n(\mathbf{B}) \\ &= \frac{1}{n} \ln \left( \sum_{k,\ell} q_{k,\ell} S_n(\mathbf{B}^{(k,\ell)}) \right) \\ &\geq \frac{1}{n} \ln \left( \sup_{k,\ell} q_{k,\ell} S_n(\mathbf{B}^{(k,\ell)}) \right) \\ &= \frac{1}{n} \sup_{k,\ell} \left( \ln q_{k,\ell} + \ln S_n(\mathbf{B}^{(k,\ell)}) \right) \end{aligned}$$

We have to prove that

$$\liminf_{n \rightarrow \infty} W_n(\mathbf{B}) = \liminf_{n \rightarrow \infty} \frac{1}{n} \ln S_n(\mathbf{B}) \geq W^* \quad \text{a.s.}$$

W.l.o.g. we may assume  $S_0 = 1$ , so that

$$\begin{aligned} W_n(\mathbf{B}) &= \frac{1}{n} \ln S_n(\mathbf{B}) \\ &= \frac{1}{n} \ln \left( \sum_{k,\ell} q_{k,\ell} S_n(\mathbf{B}^{(k,\ell)}) \right) \\ &\geq \frac{1}{n} \ln \left( \sup_{k,\ell} q_{k,\ell} S_n(\mathbf{B}^{(k,\ell)}) \right) \\ &= \frac{1}{n} \sup_{k,\ell} \left( \ln q_{k,\ell} + \ln S_n(\mathbf{B}^{(k,\ell)}) \right) \\ &= \sup_{k,\ell} \left( W_n(\mathbf{B}^{(k,\ell)}) + \frac{\ln q_{k,\ell}}{n} \right). \end{aligned}$$

Thus

$$\liminf_{n \rightarrow \infty} W_n(\mathbf{B}) \geq \liminf_{n \rightarrow \infty} \sup_{k, \ell} \left( W_n(\mathbf{B}^{(k, \ell)}) + \frac{\ln q_{k, \ell}}{n} \right)$$

Thus

$$\begin{aligned} \liminf_{n \rightarrow \infty} W_n(\mathbf{B}) &\geq \liminf_{n \rightarrow \infty} \sup_{k, \ell} \left( W_n(\mathbf{B}^{(k, \ell)}) + \frac{\ln q_{k, \ell}}{n} \right) \\ &\geq \sup_{k, \ell} \liminf_{n \rightarrow \infty} \left( W_n(\mathbf{B}^{(k, \ell)}) + \frac{\ln q_{k, \ell}}{n} \right) \end{aligned}$$

Thus

$$\begin{aligned}\liminf_{n \rightarrow \infty} W_n(\mathbf{B}) &\geq \liminf_{n \rightarrow \infty} \sup_{k, \ell} \left( W_n(\mathbf{B}^{(k, \ell)}) + \frac{\ln q_{k, \ell}}{n} \right) \\ &\geq \sup_{k, \ell} \liminf_{n \rightarrow \infty} \left( W_n(\mathbf{B}^{(k, \ell)}) + \frac{\ln q_{k, \ell}}{n} \right) \\ &= \sup_{k, \ell} \liminf_{n \rightarrow \infty} W_n(\mathbf{B}^{(k, \ell)})\end{aligned}$$

Thus

$$\begin{aligned}\liminf_{n \rightarrow \infty} W_n(\mathbf{B}) &\geq \liminf_{n \rightarrow \infty} \sup_{k, \ell} \left( W_n(\mathbf{B}^{(k, \ell)}) + \frac{\ln q_{k, \ell}}{n} \right) \\ &\geq \sup_{k, \ell} \liminf_{n \rightarrow \infty} \left( W_n(\mathbf{B}^{(k, \ell)}) + \frac{\ln q_{k, \ell}}{n} \right) \\ &= \sup_{k, \ell} \liminf_{n \rightarrow \infty} W_n(\mathbf{B}^{(k, \ell)}) \\ &= \sup_{k, \ell} \epsilon_{k, \ell}\end{aligned}$$

Since the partitions  $\mathcal{P}_\ell$  are nested, we have that

$$\sup_{k, \ell} \epsilon_{k, \ell} = \lim_{k \rightarrow \infty} \lim_{l \rightarrow \infty} \epsilon_{k, l} = W^*.$$



# Kernel regression estimate

Kernel function  $K(x) \geq 0$

Bandwidth  $h > 0$

$$m_n(x) = \frac{\sum_{i=1}^n Y_i K\left(\frac{x-X_i}{h}\right)}{\sum_{i=1}^n K\left(\frac{x-X_i}{h}\right)}$$

Naive (window) kernel function  $K(x) = I_{\{\|x\| \leq 1\}}$

$$m_n(x) = \frac{\sum_{i=1}^n Y_i I_{\{\|x-X_i\| \leq h\}}}{\sum_{i=1}^n I_{\{\|x-X_i\| \leq h\}}}$$

choose the radius  $r_{k,l} > 0$  such that for any fixed  $k$ ,

$$\lim_{l \rightarrow \infty} r_{k,l} = 0.$$

choose the radius  $r_{k,\ell} > 0$  such that for any fixed  $k$ ,

$$\lim_{\ell \rightarrow \infty} r_{k,\ell} = 0.$$

for  $n > k + 1$ , define the expert  $\mathbf{b}^{(k,\ell)}$  by

$$\mathbf{b}^{(k,\ell)}(\mathbf{x}_1^{n-1}) = \arg \max_{\mathbf{b}} \sum_{\{k < i < n: \|\mathbf{x}_{i-k}^{i-1} - \mathbf{x}_{n-k}^{n-1}\| \leq r_{k,\ell}\}} \ln \langle \mathbf{b}, \mathbf{x}_i \rangle,$$

if the sum is non-void, and  $\mathbf{b}_0 = (1/d, \dots, 1/d)$  otherwise.

The kernel-based portfolio scheme is universally consistent with respect to the class of all ergodic processes such that  $\mathbf{E}\{|\ln X^{(j)}| < \infty, \text{ for } j = 1, 2, \dots, d.$

L. Györfi, G. Lugosi, F. Udina (2006) "Nonparametric kernel-based sequential investment strategies", *Mathematical Finance*, 16, pp. 337-357

[www.szit.bme.hu/~gyorfi/kernel.pdf](http://www.szit.bme.hu/~gyorfi/kernel.pdf)

## $k$ -nearest neighbor (NN) regression estimate

$$m_n(x) = \sum_{i=1}^n W_{ni}(x; X_1, \dots, X_n) Y_i.$$

$W_{ni}$  is  $1/k$  if  $X_i$  is one of the  $k$  nearest neighbors of  $x$  among  $X_1, \dots, X_n$ , and  $W_{ni}$  is 0 otherwise.

# Nearest-neighbor-based portfolio selection

choose  $p_\ell \in (0, 1)$  such that  $\lim_{\ell \rightarrow \infty} p_\ell = 0$

for fixed positive integers  $k, \ell$  ( $n > k + \hat{\ell} + 1$ ) introduce the set of the  $\hat{\ell} = \lfloor p_\ell n \rfloor$  nearest neighbor matches:

$$\hat{J}_n^{(k, \ell)} = \{i; k + 1 \leq i \leq n \text{ such that } \mathbf{x}_{i-k}^{i-1} \text{ is among the } \hat{\ell} \text{ NNs of } \mathbf{x}_{n-k}^{n-1} \text{ in } \mathbf{x}_1^k, \dots, \mathbf{x}_{n-k}^{n-1}\}.$$

# Nearest-neighbor-based portfolio selection

choose  $p_\ell \in (0, 1)$  such that  $\lim_{\ell \rightarrow \infty} p_\ell = 0$   
for fixed positive integers  $k, \ell$  ( $n > k + \hat{\ell} + 1$ ) introduce the set of  
the  $\hat{\ell} = \lfloor p_\ell n \rfloor$  nearest neighbor matches:

$$\hat{J}_n^{(k, \ell)} = \{i; k + 1 \leq i \leq n \text{ such that } \mathbf{x}_{i-k}^{i-1} \text{ is among the } \hat{\ell} \text{ NNs of } \mathbf{x}_{n-k}^{n-1} \\ \text{in } \mathbf{x}_1^k, \dots, \mathbf{x}_{n-k}^{n-1}\}.$$

Define the portfolio vector by

$$\mathbf{b}^{(k, \ell)}(\mathbf{x}_1^{n-1}) = \arg \max_{\mathbf{b}} \sum_{\{i \in \hat{J}_n^{(k, \ell)}\}} \ln \langle \mathbf{b}, \mathbf{x}_i \rangle$$

if the sum is non-void, and  $\mathbf{b}_0 = (1/d, \dots, 1/d)$  otherwise.

If for any vector  $\mathbf{s} = \mathbf{s}_1^k$  the random variable

$$\|\mathbf{X}_1^k - \mathbf{s}\|$$

has continuous distribution function, then the nearest-neighbor portfolio scheme is universally consistent with respect to the class of all ergodic processes such that  $\mathbf{E}\{|\ln X^{(j)}|\} < \infty$ , for  $j = 1, 2, \dots, d$ .

NN is robust, there is no scaling problem



If for any vector  $\mathbf{s} = \mathbf{s}_1^k$  the random variable

$$\|\mathbf{X}_1^k - \mathbf{s}\|$$

has continuous distribution function, then the nearest-neighbor portfolio scheme is universally consistent with respect to the class of all ergodic processes such that  $\mathbf{E}\{|\ln X^{(j)}|\} < \infty$ , for  $j = 1, 2, \dots, d$ .

NN is robust, there is no scaling problem

L. Györfi, F. Udina, H. Walk (2006) "Nonparametric nearest neighbor based empirical portfolio selection strategies", (submitted), [www.szit.bme.hu/~gyorfi/NN.pdf](http://www.szit.bme.hu/~gyorfi/NN.pdf)

# Semi-log-optimal portfolio

empirical log-optimal:

$$h^{(k,\ell)}(\mathbf{x}_1^{n-1}) = \arg \max_{\mathbf{b}} \sum_{i \in J_n} \ln \langle \mathbf{b}, \mathbf{x}_i \rangle$$

# Semi-log-optimal portfolio

empirical log-optimal:

$$h^{(k,\ell)}(\mathbf{x}_1^{n-1}) = \arg \max_{\mathbf{b}} \sum_{i \in J_n} \ln \langle \mathbf{b}, \mathbf{x}_i \rangle$$

Taylor expansion:  $\ln z \approx h(z) = z - 1 - \frac{1}{2}(z - 1)^2$

# Semi-log-optimal portfolio

empirical log-optimal:

$$h^{(k,\ell)}(\mathbf{x}_1^{n-1}) = \arg \max_{\mathbf{b}} \sum_{i \in J_n} \ln \langle \mathbf{b}, \mathbf{x}_i \rangle$$

Taylor expansion:  $\ln z \approx h(z) = z - 1 - \frac{1}{2}(z - 1)^2$  empirical  
semi-log-optimal:

$$\tilde{h}^{(k,\ell)}(\mathbf{x}_1^{n-1}) = \arg \max_{\mathbf{b}} \sum_{i \in J_n} h(\langle \mathbf{b}, \mathbf{x}_i \rangle) = \arg \max_{\mathbf{b}} \{ \langle \mathbf{b}, \mathbf{m} \rangle - \langle \mathbf{b}, \mathbf{C}\mathbf{b} \rangle \}$$

# Semi-log-optimal portfolio

empirical log-optimal:

$$h^{(k,\ell)}(\mathbf{x}_1^{n-1}) = \arg \max_{\mathbf{b}} \sum_{i \in J_n} \ln \langle \mathbf{b}, \mathbf{x}_i \rangle$$

Taylor expansion:  $\ln z \approx h(z) = z - 1 - \frac{1}{2}(z - 1)^2$  empirical  
semi-log-optimal:

$$\tilde{h}^{(k,\ell)}(\mathbf{x}_1^{n-1}) = \arg \max_{\mathbf{b}} \sum_{i \in J_n} h(\langle \mathbf{b}, \mathbf{x}_i \rangle) = \arg \max_{\mathbf{b}} \{ \langle \mathbf{b}, \mathbf{m} \rangle - \langle \mathbf{b}, \mathbf{C}\mathbf{b} \rangle \}$$

smaller computational complexity: quadratic programming

empirical log-optimal:

$$h^{(k,\ell)}(\mathbf{x}_1^{n-1}) = \arg \max_{\mathbf{b}} \sum_{i \in J_n} \ln \langle \mathbf{b}, \mathbf{x}_i \rangle$$

Taylor expansion:  $\ln z \approx h(z) = z - 1 - \frac{1}{2}(z - 1)^2$  empirical semi-log-optimal:

$$\tilde{h}^{(k,\ell)}(\mathbf{x}_1^{n-1}) = \arg \max_{\mathbf{b}} \sum_{i \in J_n} h(\langle \mathbf{b}, \mathbf{x}_i \rangle) = \arg \max_{\mathbf{b}} \{ \langle \mathbf{b}, \mathbf{m} \rangle - \langle \mathbf{b}, \mathbf{C}\mathbf{b} \rangle \}$$

smaller computational complexity: quadratic programming

L. Györfi, A. Urbán, I. Vajda (2007) "Kernel-based semi-log-optimal portfolio selection strategies", *International Journal of Theoretical and Applied Finance*, 10, pp. 505-516.  
[www.szit.bme.hu/~gyorfi/semi.pdf](http://www.szit.bme.hu/~gyorfi/semi.pdf)

# Conditions of the model:

Assume that

- the assets are arbitrarily divisible,
- the assets are available in unbounded quantities at the current price at any given trading period,
- there are no transaction costs,
- the behavior of the market is not affected by the actions of the investor using the strategy under investigation.

At [www.szit.bme.hu/~oti/portfolio](http://www.szit.bme.hu/~oti/portfolio) there are two benchmark data set from NYSE:

- The first data set consists of daily data of 36 stocks with length 22 years.
- The second data set contains 23 stocks and has length 44 years.



At [www.szit.bme.hu/~oti/portfolio](http://www.szit.bme.hu/~oti/portfolio) there are two benchmark data set from NYSE:

- The first data set consists of daily data of 36 stocks with length 22 years.
- The second data set contains 23 stocks and has length 44 years.

Our experiment is on the second data set.

# Experiments on average annual yields (AAY)

Kernel based semi-log-optimal portfolio selection with  
 $k = 1, \dots, 5$  and  $l = 1, \dots, 10$

$$r_{k,l}^2 = 0.0001 \cdot d \cdot k \cdot l,$$

# Experiments on average annual yields (AAY)

Kernel based semi-log-optimal portfolio selection with  
 $k = 1, \dots, 5$  and  $l = 1, \dots, 10$

$$r_{k,l}^2 = 0.0001 \cdot d \cdot k \cdot l,$$

AAY of kernel based semi-log-optimal portfolio is 128%

# Experiments on average annual yields (AAY)

Kernel based semi-log-optimal portfolio selection with  
 $k = 1, \dots, 5$  and  $l = 1, \dots, 10$

$$r_{k,l}^2 = 0.0001 \cdot d \cdot k \cdot l,$$

AAY of kernel based semi-log-optimal portfolio is 128%  
double the capital

# Experiments on average annual yields (AAY)

Kernel based semi-log-optimal portfolio selection with  
 $k = 1, \dots, 5$  and  $l = 1, \dots, 10$

$$r_{k,l}^2 = 0.0001 \cdot d \cdot k \cdot l,$$

AAY of kernel based semi-log-optimal portfolio is 128%  
double the capital  
MORRIS had the best AAY, 20%

# Experiments on average annual yields (AAY)

Kernel based semi-log-optimal portfolio selection with  
 $k = 1, \dots, 5$  and  $l = 1, \dots, 10$

$$r_{k,l}^2 = 0.0001 \cdot d \cdot k \cdot l,$$

AAY of kernel based semi-log-optimal portfolio is 128%  
double the capital

MORRIS had the best AAY, 20%

the BCRP had average AAY 24%

# The average annual yields of the individual experts.

$k$	1	2	3	4	5
$\ell$					
1	20%	19%	16%	16%	16%
2	118%	77%	62%	24%	58%
3	71%	41%	26%	58%	21%
4	103%	94%	63%	97%	34%
5	134%	102%	100%	102%	67%
6	140%	125%	105%	108%	87%
7	148%	123%	107%	99%	96%
8	132%	112%	102%	85%	81%
9	127%	103%	98%	74%	72%
10	123%	92%	81%	65%	69%